

# Brain beta-amyloid measures and magnetic resonance imaging atrophy both predict time-to-progression from mild cognitive impairment to Alzheimer's disease

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Biomarkers of brain A $\beta$  amyloid deposition can be measured either by cerebrospinal fluid A $\beta$ 42 or Pittsburgh compound B positron emission tomography imaging. Our objective was to evaluate the ability of A $\beta$  load and neurodegenerative atrophy on magnetic resonance imaging to predict shorter time-to-progression from mild cognitive impairment to Alzheimer's dementia and to characterize the effect of these biomarkers on the risk of progression as they become increasingly abnormal. A total of 218 subjects with mild cognitive impairment were identified from the Alzheimer's Disease Neuroimaging Initiative. The primary outcome was time-to-progression to Alzheimer's dementia. Hippocampal volumes were measured and adjusted for intracranial volume. We used a new method of pooling cerebrospinal fluid A $\beta$ 42 and Pittsburgh compound B positron emission tomography measures to produce equivalent measures of brain A $\beta$  load from either source and analysed the results using multiple imputation methods. We performed our analyses in two phases. First, we grouped our subjects into those who were 'amyloid positive' ( $n=165$ , with the assumption that Alzheimer's pathology is dominant in this group) and those who were 'amyloid negative' ( $n=53$ ). In the second phase, we included all 218 subjects with mild cognitive impairment to evaluate the biomarkers in a sample that we assumed to contain a full spectrum of expected pathologies. In a Kaplan–Meier analysis, amyloid positive subjects with mild cognitive impairment were much more likely to progress to dementia within 2 years than amyloid negative subjects with mild cognitive impairment (50 versus 19%). Among amyloid positive subjects with mild cognitive impairment

only, hippocampal atrophy predicted shorter time-to-progression ( $P < 0.001$ ) while A $\beta$  load did not ( $P = 0.44$ ). In contrast, when all 218 subjects with mild cognitive impairment were combined (amyloid positive and negative), hippocampal atrophy and A $\beta$  load predicted shorter time-to-progression with comparable power (hazard ratio for an inter-quartile difference of 2.6 for both); however, the risk profile was linear throughout the range of hippocampal atrophy values but reached a ceiling at higher values of brain A $\beta$  load. Our results are consistent with a model of Alzheimer's disease in which A $\beta$  deposition initiates the pathological cascade but is not the direct cause of cognitive impairment as evidenced by the fact that A $\beta$  load severity is decoupled from risk of progression at high levels. In contrast, hippocampal atrophy indicates how far along the neurodegenerative path one is, and hence how close to progressing to dementia. Possible explanations for our finding that many subjects with mild cognitive impairment have intermediate levels of A $\beta$  load include: (i) individual subjects may reach an A $\beta$  load plateau at varying absolute levels; (ii) some subjects may be more biologically susceptible to A $\beta$  than others; and (iii) subjects with mild cognitive impairment with intermediate levels of A $\beta$  may represent individuals with Alzheimer's disease co-existent with other pathologies.

**Keywords:** mild cognitive impairment; amyloid imaging; magnetic resonance imaging; cerebrospinal fluid; Alzheimer's disease biomarkers

**Abbreviation:** PIB = Pittsburgh compound B

## Introduction

The most widely accepted and validated biomarkers in Alzheimer's disease fall into two categories: imaging and CSF chemical analytes (Shaw *et al.*, 2007; Hampel *et al.*, 2008). Different biomarkers serve as *in vivo* indicators of specific pathologies. Measures of brain atrophy on MRI are biomarkers of neurodegenerative pathology (Bobinski *et al.*, 2000; Gosche *et al.*, 2002; Jack *et al.*, 2002; Silbert *et al.*, 2003; Jagust *et al.*, 2008; Vemuri *et al.*, 2008b; Whitwell *et al.*, 2008), while both PET amyloid imaging (Klunk *et al.*, 2004; Edison *et al.*, 2007; Rowe *et al.*, 2007; Drzezga *et al.*, 2008; Ikonovic *et al.*, 2008; Leinonen *et al.*, 2008; Frisoni *et al.*, 2009; Tolboom *et al.*, 2009) and decreased CSF A $\beta$ 42 (Clark *et al.*, 2003; Strozyk *et al.*, 2003; Schoonenboom *et al.*, 2008; Buchhave *et al.*, 2009; Tapiola *et al.*, 2009) are indicators of brain A $\beta$  amyloidosis (referred to from here on as A $\beta$  load).

One of the most meaningful clinical applications of Alzheimer's disease biomarkers is as an aid to predicting future clinical course. Measures of brain atrophy on MRI are well-established predictors of progression from mild cognitive impairment to Alzheimer's disease (Jack *et al.*, 1999, 2005; Visser *et al.*, 1999; Killiany *et al.*, 2000; Chetelat *et al.*, 2005; Stoub *et al.*, 2005; Apostolova *et al.*, 2006; Devanand *et al.*, 2007; Vemuri *et al.*, 2008a; Davatzikos *et al.*, 2009; Dickerson *et al.*, 2009; Driscoll *et al.*, 2009; Fennema-Notestine *et al.*, 2009; McEvoy *et al.*, 2009; Risacher *et al.*, 2009). The presence of significant brain A $\beta$  load, measured either by CSF A $\beta$ 42 or PET amyloid imaging, is also highly correlated with progression from mild cognitive impairment to Alzheimer's disease (Hansson *et al.*, 2006; Forsberg *et al.*, 2008; Brys *et al.*, 2009; Mattsson *et al.*, 2009; Okello *et al.*, 2009; Visser *et al.*, 2009; Waragai *et al.*, 2009; Wolk *et al.*, 2009). However, evidence indicates that A $\beta$  accumulation begins as much as decades prior to the appearance of the first cognitive symptoms (Mintun *et al.*, 2006; Peskind *et al.*, 2006; Aizenstein *et al.*, 2008; Bouwman *et al.*, 2009; Kok *et al.*, 2009; Reiman

*et al.*, 2009; Scheinin *et al.*, 2009; Sperling *et al.*, 2009; Bourgeat *et al.*, 2010;), while anatomical MRI becomes abnormal later in the disease course (Fox *et al.*, 1996, 2001; Carlson *et al.*, 2008). These findings suggest that the associations between abnormalities in these two classes of biomarkers and the 'time-dependent risk' of progressing from mild cognitive impairment to Alzheimer's disease may differ, and this has not yet been investigated to our knowledge.

We used a new method of transforming CSF A $\beta$ 42 measures into units of Pittsburgh compound B (PIB) PET (Weigand *et al.*, 2010) and pooled data from patients who had only one or the other measure of A $\beta$  load using multiple imputation measurement error models (Cole *et al.*, 2006). Our objectives were to: (i) measure the ability of biomarkers of A $\beta$  load and neurodegeneration (using MRI) to predict shorter time-to-progression from mild cognitive impairment to Alzheimer's disease; and (ii) determine how the severity of these two categories of biomarker affects the time-to-progression by estimating the log relative hazard as a possible non-linear function of biomarker severity. To incorporate the aetiological heterogeneity of subjects who meet clinical criteria for mild cognitive impairment in our discussion, we performed our analyses in two phases. In the first, we grouped our subjects into those who were 'amyloid positive', in whom we assume Alzheimer's disease is the dominant pathology, and those who were 'amyloid negative'. In the second phase, we included all subjects with mild cognitive impairment in order to compare MRI and A $\beta$  load biomarkers as predictors in a cohort that included the full spectrum of expected pathologies associated with the clinical syndrome of mild cognitive impairment.

## Materials and methods

### Subjects

A total of 218 subjects with a diagnosis of mild cognitive impairment (Petersen *et al.*, 2001) and one or more clinical follow-up assessments

were identified from the Alzheimer's Disease Neuroimaging Initiative (ADNI) (Petersen *et al.*, 2010). Subjects must have undergone either lumbar puncture or PIB PET while carrying a diagnosis of mild cognitive impairment. The primary outcome was time-to-progression from a clinical diagnosis of mild cognitive impairment to Alzheimer's disease. The diagnosis of dementia was made using DSM-IV criteria (1994), and the diagnosis of Alzheimer's disease was made using established clinical criteria (McKhann *et al.*, 1984).

## Magnetic resonance imaging methods

All subjects were scanned at 1.5 T with a 3D magnetization prepared rapid acquisition gradient echo imaging sequence developed at the Mayo Clinic Rochester for the Alzheimer's Disease Neuroimaging Initiative study (Jack *et al.*, 2008). All images were corrected for image distortion due to gradient non-linearity using 'GradWarp' (Jovicich *et al.*, 2006), correction for B1 non-uniformity as necessary (Jack *et al.*, 2008) and for residual inhomogeneity using 'N3' (Sled *et al.*, 1998) with a software pipeline running at the Mayo Clinic Rochester. Hippocampal volumes and total intracranial volumes were measured at Mayo Clinic using FreeSurfer software (version 4.5.0) (Fischl *et al.*, 2002). Hippocampal volumes were adjusted for total intracranial volumes by including total intracranial volumes as a covariate in the Cox models (Jack *et al.*, 1989).

## Amyloid imaging methods

PIB PET studies were performed at 13 different sites. Production of PIB PET and radio labelling with  $^{11}\text{C}$  was performed as outlined in Mathis *et al.* (2003). The PIB PET images undergo several quality control and standardization steps, which are described at <http://www.ADNI-info.org>. The PIB PET images used in our study were the 'maximally pre-processed files' available for download.

All PIB PET quantitative image analysis was performed at the Mayo Clinic using the same fully automated image processing pipeline as described in Senjem *et al.* (2008; Jack, 2008). The method includes partial volume correction and region of interest sharpening of PIB PET images using each subject's spatially registered MRI. Statistics on image voxel values were extracted from automatically labelled cortical regions of interest using an in-house modification of the automated anatomic labelling atlas (Tzourio-Mazoyer *et al.*, 2002). A global cortical PIB PET retention summary was formed by combining the prefrontal, orbitofrontal, parietal, temporal, anterior cingulate and posterior cingulate/precuneus values for each subject, using a weighted average of these regions of interest values where larger regions of interest were given greater weight. PIB PET ratio values were calculated by dividing the median value in each target cortical region of interest by the median value in the cerebellar grey matter region of interest of the atlas.

## Cerebrospinal fluid methods

CSF was collected at each site, transferred into polypropylene transfer tubes followed by freezing on dry ice within 1 h after collection and shipped overnight to the Alzheimer's Disease Neuroimaging Initiative Biomarker Core laboratory at the University of Pennsylvania Medical Centre on dry ice. A standardized protocol was implemented to quantify biomarker concentrations in each of the CSF baseline aliquots using a multiplex xMAP Luminex platform (Luminex Corp, Austin, TX) with Innogenetics (INNO-BIA AlzBio3, Ghent, Belgium; for research use only reagents) immunoassay kit-based reagents, validated

in Vanderstichele *et al.* (2008) and Shaw *et al.* (2009). Details can be found at (<http://www.adni-info.org/index.php>).

## Statistical methods

We used Cox proportional hazards models to estimate the effect of A $\beta$  load and hippocampal volume on the relative hazard of progression from mild cognitive impairment to Alzheimer's disease. In Cox models, increased relative hazard is directly related to shortened time-to-event. Time 0 for each subject was defined as the date of their earliest visit with PIB PET imaging, or with CSF if the subject did not participate in PIB PET imaging. The event time was defined as the midpoint between the last visit in which the subject was diagnosed with mild cognitive impairment and the first visit in which the subject was diagnosed as demented. Subjects who were never observed to progress to dementia were censored at their last visit. Two subjects who met the inclusion criteria but progressed to a clinically diagnosed non-Alzheimer's dementia were censored at their event time because at this point they were no longer at risk for Alzheimer's dementia and in a time-to-event analysis it would be inappropriate to remove these subjects since they met the baseline inclusion criteria.

The primary predictors of interest were hippocampal volume with total intracranial volumes included as a covariate and A $\beta$  load. To allow for a non-linear relationship between the predictors and the log hazard, we modelled the predictors using restricted cubic splines with knots at the 10th, 50th and 90th percentiles (Harrell, 2001). We report hazard ratios based on comparing the 25th, 50th and 75th percentiles to aid interpretability across biomarkers. The 25th percentile can be thought of as a typical 'low' value since it is the middle value among those below the median (50th percentile). Similarly, the 75th percentile can be thought of as a typical 'high' value. Since higher values of A $\beta$  are more abnormal while lower values of hippocampal volume are more abnormal, a hazard ratio comparing the 75th to the 25th percentile for A $\beta$  load is analogous to one comparing the 25th to the 75th percentile for hippocampal volume.

Throughout the manuscript, measures of brain A $\beta$  load are expressed in the 'cortical-to-cerebellar ratio units', which are typically used to measure PIB PET retention and range from 1.0 to ~3.0. These values are referred to as A $\beta$  load whether derived from CSF A $\beta$ 42 or actual PIB PET imaging. We used a multiple imputation measurement error approach to transform CSF A $\beta$ 42 into PIB PET units and pool measures of A $\beta$  from either source (Cole *et al.*, 2006). We describe the necessary steps in detail elsewhere but briefly summarize the method here (Weigand *et al.*, 2010). A calibration data set of 41 subjects who participated in PIB PET imaging at the time of their lumbar puncture was used to estimate the relationship between PIB PET retention ( $y$ ), CSF A $\beta$ 42 ( $x$ ) and whether the subject carries the APOE  $\epsilon$ 4 allele ( $z$ ). The fitted linear regression 'conversion model' was found to be:  $\log_2(y) = 5.326 - 0.615 \log_2(x) + 0.184(z) + e$ , where  $y$  is the estimated PIB PET value that we call PIBcalc,  $x$  is the CSF-based value,  $z$  is 0 if the subject carries no APOE  $\epsilon$ 4 alleles and 1 if the subject is an  $\epsilon$ 4 carrier, and  $e$  is a random error term that is normally distributed with mean 0 and SD 0.180. While this formula can be used to obtain a 'best guess' estimate of a subject's PIB PET retention, treating the result as if it were obtained directly from PIB PET imaging is inappropriate because it ignores the error term in the conversion model and the uncertainty associated with the conversion model coefficients, which are estimated rather than known exactly. To correctly carry forward prediction uncertainty and model estimation uncertainty to subsequent stages of the analysis, the multiple imputation measurement error approach uses multiple imputation (Little and Rubin, 2002).

We generated 100 multiple imputation data sets where in each data set a subject's A $\beta$  load value was the PIB PET value if available or a simulated PIBcalc value otherwise. These simulated values incorporate the error term from the conversion model plus an additional perturbation to account for uncertainty in the conversion model parameters. We then fit a Cox proportional hazard model as described above to each data set and pooled the results using the combining rules of multiple imputation. We used multiple imputation likelihood ratio tests to perform multiple degree of freedom tests of the linearity of a biomarker predictor (Harel and Zhou, 2007). To compare progression among those above versus below an A $\beta$  load level of 1.5, we performed a Kaplan–Meier analysis on each multiple imputation data set by dichotomizing the amyloid load variable at 1.5. Data manipulation was performed using SAS version 9.1.3 (SAS Institute Inc., 2004) and analysis was performed using R version 2.7.1. (R Development Core Team, 2008). We used version 2.35-9 of the survival package for R (Therneau, 2008) and version 2.0 of the mitools package for R (Lumley, 2008).

## Results

Analyses were performed in two phases. In the first phase, we used a generally accepted cut-off value reported in the PIB PET literature of 1.5 to classify subjects as either 'amyloid positive' versus 'amyloid negative' (Rowe *et al.*, 2007). Amyloid positive subjects ( $n=165$ ) were more frequently APOE  $\epsilon 4$  carriers, performed slightly worse on the Clinical Dementia Rating Scale – Sum of Boxes (CDR-SB) and had smaller hippocampal volumes at baseline than amyloid negative subjects ( $n=53$ ) (Table 1). Amyloid positive subjects had a 3-fold increase in hazard of progressing (hazard ratio 3.2, 95% CI 1.4–7.1,  $P=0.004$ ) with Kaplan–Meier analysis indicating 50% of amyloid positive subjects will progress to dementia in 2 years versus 19% among amyloid negative. Figure 1 illustrates MRI and PIB PET imaging findings in a typical progressor and a typical non-progressor. With only eight progressors among the amyloid negative group, relationships between progression and biomarkers could not reliably be assessed.

Among the amyloid positive subjects, the hazard ratio (95% CI) for progressing was 1.2 (0.9–1.8) ( $P=0.44$ ) for an individual in the 75th versus 25th percentile of the A $\beta$  load distribution (whose lower bound was by definition 1.5). The analogous hazard ratio (95% CI) for the 25th versus 75th percentile of hippocampal volume was 2.0 (1.5–2.8) ( $P<0.001$ ). The relationship between log relative hazard of progressing and increasingly atrophic hippocampi was essentially linear ( $P=0.93$ , versus a non-linear spline fit).

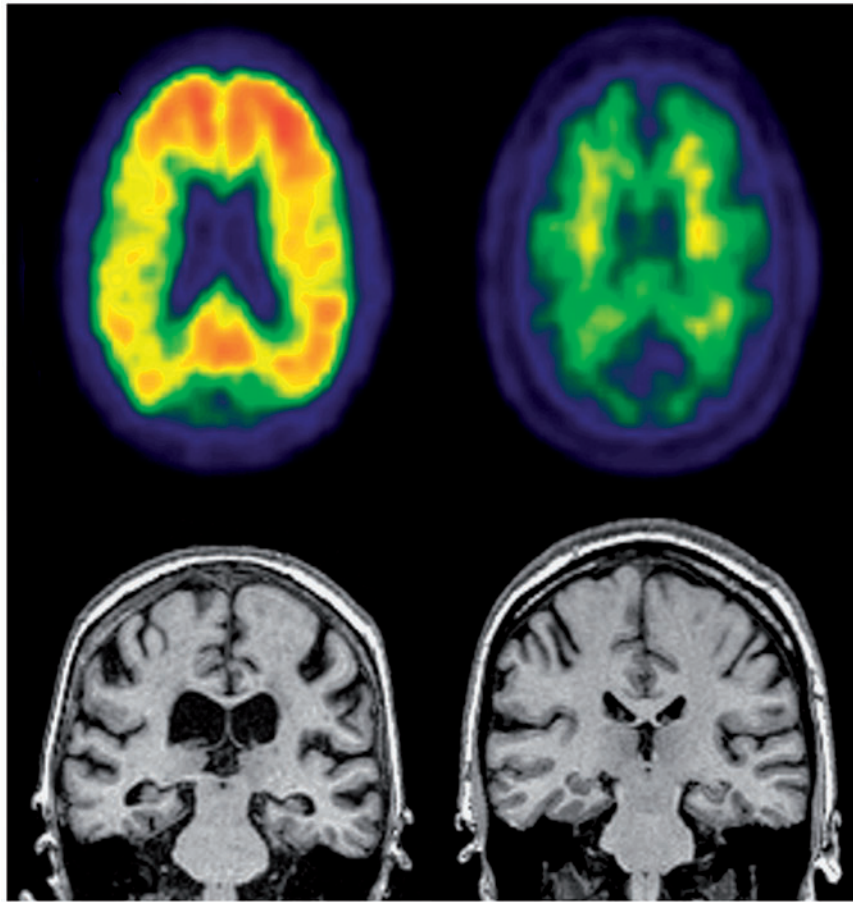
Among all 218 subjects with mild cognitive impairment combined, over a median progression-free follow-up time of 1.7 years, 89 subjects progressed from mild cognitive impairment to dementia. In qualitative terms, age and education did not differ between progressors and non-progressors, although women made up a larger proportion of the progressors (Table 1). The progressor group had a higher proportion of APOE  $\epsilon 4$  carriers, and slightly worse scores on the Mini-Mental State Examination and Clinical Dementia Rating Scale B than non-progressors at baseline. Progressors had greater A $\beta$  load and more atrophic hippocampi at baseline than non-progressors. Direct statistical comparisons of progressor versus non-progressor are inappropriate because follow-up times varied across individual subjects; hence our use of time-to-event statistical methods with results reported as hazard ratios. Among all subjects with mild cognitive impairment the hazard ratio (95% CI) for progressing was 2.6 (1.5–4.5) for an individual in the 75th versus 25th percentile of the A $\beta$  load distribution. The analogous hazard ratio (95% CI) for the 25th versus 75th percentile of hippocampal volume was 2.6 (1.8–3.8). (The percentiles being compared are reversed because greater A $\beta$  is associated with greater risk, while smaller hippocampal volume is associated with greater risk.) Both A $\beta$  load and MRI were highly significant predictors of progression overall ( $P\leq 0.001$  for both) (Table 2). The relationship between log relative hazard of progressing and increasingly atrophic hippocampi was essentially linear ( $P=0.60$  versus a non-linear spline fit) (Fig. 2). In contrast there was evidence of non-linearity ( $P=0.060$ ) in this relationship for A $\beta$  load—such that a ceiling was reached in the log relative

**Table 1** Descriptive characteristics of all 218 subjects with mild cognitive impairment by progressor status, and by A $\beta$  load status

Characteristic	All	Stable <sup>a</sup>	Progressor	Amyloid negative ( $\leq 1.5$ )	Amyloid positive ( $> 1.5$ )
Number of subjects	$n=218$	$n=129$	$n=89$	$n=53$	$n=165$
Female gender, number (%)	72 (33)	38 (29)	34 (38)	13 (25)	59 (36)
Age, years, median (interquartile range)	75 (70, 80)	75 (70, 81)	75 (70, 80)	77 (70, 83)	75 (70, 80)
APOE positive, number (%)	117 (54)	58 (45)	59 (66)	7 (13)	110 (67)
Education, years, median (interquartile range)	16 (14, 18)	16 (14, 18)	16 (14, 18)	16 (14, 18)	16 (14, 18)
MMSE, median (interquartile range)	27 (25, 29)	28 (26, 29)	26 (25, 28)	28 (26, 29)	27 (25, 28)
CDR-SB, median (interquartile range)	1.5 (1.0, 2.0)	1.0 (1.0, 1.5)	2.0 (1.0, 2.5)	1.0 (1.0, 1.5)	1.5 (1.0, 2.0)
Hippocampal volume, cm <sup>3</sup> , median (interquartile range)	6.3 (5.6, 7.1)	6.7 (6.0, 7.5)	5.9 (5.0, 6.5)	6.9 (5.6, 7.6)	6.2 (5.6, 6.8)
A $\beta$ , median (interquartile range)	2.0 (1.5, 2.3)	1.8 (1.4, 2.2)	2.2 (1.9, 2.3)	1.3 (1.3, 1.4)	2.2 (1.9, 2.4)
Number (%) who were amyloid 'positive'	165 (76)	84 (65)	81 (91)	0 (0)	165 (100)
Number (%) in whom A $\beta$ load was measured by PIB PET	53 (24)	35 (27)	18 (20)	19 (36)	34 (21)

CDR-SB = Clinical Dementia Rating-Sum of Boxes ; MMSE = Mini-Mental State Examination.

<sup>a</sup> Subjects remained stable through last follow-up at which point they were censored.



**Figure 1** Illustrative images. *Left:* Mild cognitive impairment progressor, *Top:* positive PIB PET. *Bottom:* MRI illustrating atrophic hippocampi and ventricular enlargement. *Right:* Mild cognitive impairment non-progressor. *Top:* negative PIB PET with non-specific white matter retention but no cortical retention. *Bottom:* MRI illustrating normal hippocampi and no ventricular enlargement.

**Table 2** Summary of hazard ratios from Cox proportional hazard models within all 218 subjects with mild cognitive impairment

	Hazard ratio (95% CI)	P
Hippocampal volume <sup>a</sup>		<0.001
25th versus 50th percentile (i.e. 5.6 versus 1.6)	1.6 (1.3, 1.8)	
50th versus 75th percentile (i.e. 6.3 versus 7.1)	1.7 (1.2, 2.2)	
25th versus 75th percentile (i.e. 5.6 versus 7.1)	2.6 (1.8, 3.8)	
A $\beta$ load		<0.001
50th versus 25th percentile (i.e. 2.0 versus 1.5)	2.3 (1.4, 3.8)	
75th versus 50th percentile (i.e. 2.3 versus 2.0)	1.1 (0.9, 1.4)	
75th versus 25th percentile (i.e. 2.3 versus 1.5)	2.6 (1.5, 4.5)	

<sup>a</sup> Hippocampal volume model also includes total intracranial volumes as a covariate.

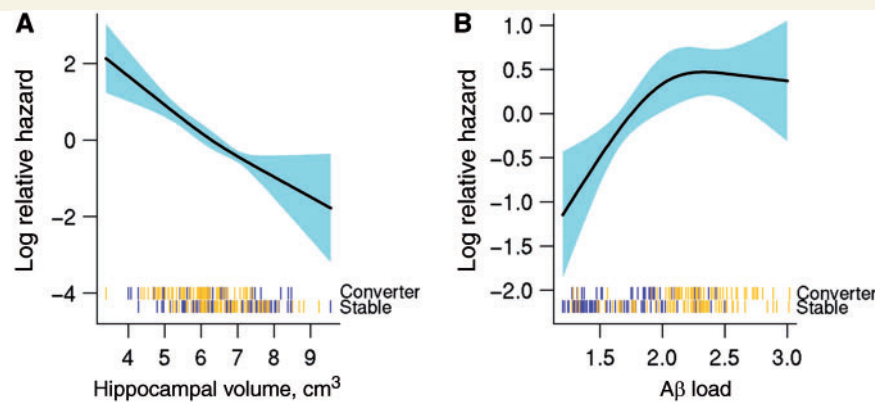
hazard as A $\beta$  load exceeded a value of  $\sim$ 2.0 (Fig. 2). The effect of severity on the hazard of progression is illustrated in the plotted log relative hazard profiles in Fig. 2 and summarized quantitatively in Table 2. The 25th versus 50th percentile hazard ratio (95% CI) for hippocampal volume was 1.6 (1.3–1.8) while the 50th versus

75th percentile hazard ratio (95% CI) was similar at 1.7 (1.2–2.2). In contrast, the 50th versus 25th percentile hazard ratio for A $\beta$  (2.3, 95% CI 1.4–3.8) was twice as great as the 75th versus 50th percentile hazard ratio (1.1, 95% CI 0.9–1.4).

Figure 2 illustrates relationships among biomarker levels, hazard of progressing and also APOE genotype. Tick marks at the bottom of each graph show the imaging data for individual progressor versus stable subjects. APOE  $\epsilon$ 4 carriers are represented with gold tick marks and non-carriers are represented with blue tick marks. A $\beta$  load and APOE  $\epsilon$ 4 status are closely related, with carriers more likely to have higher A $\beta$  load than non-carriers. In contrast, a much weaker association is seen between APOE  $\epsilon$ 4 and hippocampal atrophy.

## Discussion

Our major findings were the following: (i) amyloid positive subjects with mild cognitive impairment were much more likely to progress to a clinical diagnosis of Alzheimer's disease than amyloid negative subjects with mild cognitive impairment (50 versus 19% by 2 years); (ii) among only amyloid positive



**Figure 2** Risk profile as a function of increasing biomarker severity among all 218 subjects with mild cognitive impairment. Log hazard of progressing to dementia as a function of (A) increasing hippocampal volume (adjusting for total intracranial volumes) and (B) increasing brain A $\beta$  amyloid load within all 218 subjects with mild cognitive impairment. Hash marks at the bottom of the plot indicate the hippocampal volume and A $\beta$  load measures of individual subjects with mild cognitive impairment, with APOE genotype ( $\epsilon$ 4 carrier gold, non-carrier blue) and progressor versus non-progressor status indicated.

subjects with mild cognitive impairment, hippocampal atrophy predicted shorter time-to-progression ( $P < 0.001$ ) while amyloid load did not ( $P = 0.44$ ); (iii) In contrast, when all 218 subjects with mild cognitive impairment were combined (amyloid positive and negative), hippocampal atrophy and brain A $\beta$  load predicted time-to-progression with comparable power; and (iv) however, among all subjects with mild cognitive impairment combined, the effects of these two classes of biomarkers differ. The risk profile is linear throughout the range of hippocampal atrophy values whereas the profile reaches a ceiling at higher values of brain A $\beta$  load.

### Amyloid positive subjects with mild cognitive impairment are more likely to progress to a clinical diagnosis of Alzheimer's disease in short-term follow-up than amyloid negative subjects with mild cognitive impairment

Our results on A $\beta$  load predicting progression are consistent with studies indicating that abnormally low CSF A $\beta$ 42 is associated with an elevated risk of progressing from mild cognitive impairment to Alzheimer's disease (Hansson *et al.*, 2006; Kester *et al.*, 2009; Mattsson *et al.*, 2009) or an approximate diagnostic equivalent (Snider *et al.*, 2009). Our results are also consistent with several recent papers showing that individuals diagnosed with mild cognitive impairment who have positive PIB PET scans are more likely to progress to Alzheimer's disease in short-term follow-up than are subjects with mild cognitive impairment with negative PIB PET scans (Forsberg *et al.*, 2008; Okello *et al.*, 2009; Volk *et al.*, 2009). Although we did not examine cognitively normal elderly subjects, recent reports indicate that both abnormal PIB PET scans and low CSF A $\beta$ 42 are associated with cognitive decline in cognitively normal subjects (Skoog *et al.*, 2003; Fagan *et al.*,

2007, 2009; Gustafson *et al.*, 2007; Li *et al.*, 2007; Stomrud *et al.*, 2007; Villemagne *et al.*, 2008; Lambert *et al.*, 2009; Morris *et al.*, 2009; Chetelat *et al.*, 2010; Resnick *et al.*, 2010).

### Among only amyloid positive subjects with mild cognitive impairment, hippocampal atrophy predicts shorter time-to-progression from mild cognitive impairment to Alzheimer's disease while amyloid load does not

Evidence from multiple sources overwhelmingly points to A $\beta$  as the initiating molecular pathway in Alzheimer's disease pathogenesis. Many in the field believe that mildly impaired subjects with biomarker evidence of brain A $\beta$  amyloidosis can be presumed to have early Alzheimer's disease (Dubois *et al.*, 2007; Morris *et al.*, 2009). In contrast, mildly impaired or demented individuals who have negative A $\beta$  amyloid biomarker studies may be presumed to have non-Alzheimer's disease pathogenic substrates (Rabinovici *et al.*, 2007). Thus subjects with mild cognitive impairment who are amyloid positive can be treated as qualitatively different from those who are amyloid negative on conceptual grounds. Our intent in the amyloid positive subset analysis was to compare MRI and A $\beta$  load in those subjects with mild cognitive impairment who probably have Alzheimer's disease as the dominant pathology. An obvious criticism of this analysis is that we have restricted the range of A $\beta$  load values, thus handicapping A $\beta$  measures relative to MRI. Hippocampal atrophy is not specific for Alzheimer's disease as it occurs in other degenerative conditions (Jack *et al.*, 2002), however, we can assume that in a cognitively impaired subject with 'pure Alzheimer's disease pathology' the hippocampal atrophy observed is largely due to the Alzheimer's disease pathological process. By including the full range of hippocampal values in this subset analysis, we have

simply included the full range of neurodegenerative atrophy due to the Alzheimer's disease pathological process.

A recently described model of the Alzheimer's disease pathological cascade posits that the features of this cascade that are detectable by well-established disease biomarkers begin with detection of A $\beta$  deposition (Jack *et al.*, 2009, 2010). This model (referred to from here forward as a 'biomarker cascade model') rests on the concept that biomarker abnormalities and clinical expression of disease change over time in a sequential manner. The disease process is initiated with A $\beta$  deposition and substantial A $\beta$  deposition occurs while subjects are still cognitively normal. However, A $\beta$  amyloidosis, while necessary, is not sufficient to cause Alzheimer's dementia. A $\beta$  amyloidosis triggers a downstream neurodegenerative process that in turn is the proximate cause of cognitive impairment that progresses to dementia (DeKosky and Scheff, 1990; Terry *et al.*, 1991; Bennett *et al.*, 2005b; Jack *et al.*, 2009; Mormino *et al.*, 2009; Savva *et al.*, 2009).

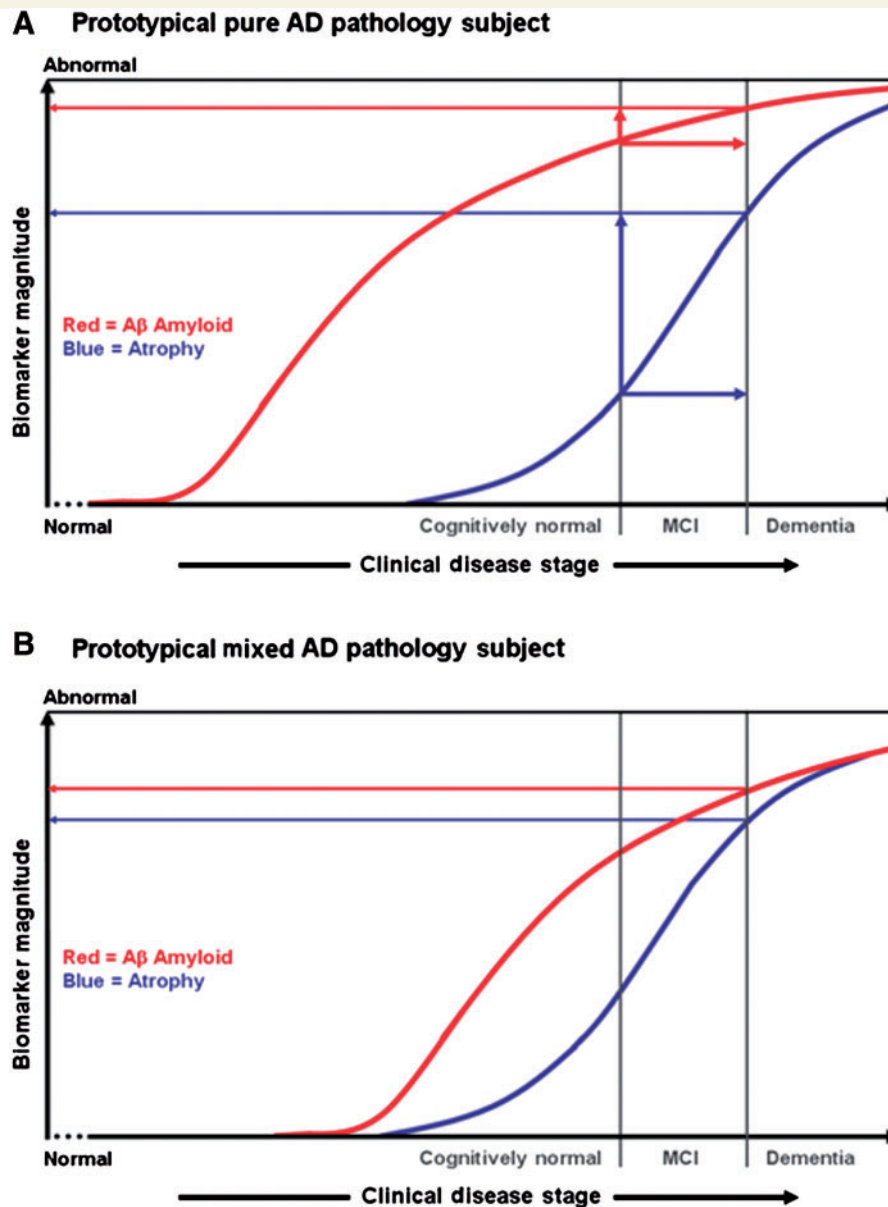
This 'biomarker cascade model' (Jack *et al.*, 2010) implies that the development of dementia due to Alzheimer's disease (as a dichotomous yes/no event over the life time of any subject) should be predicted by the aetiologically specific A $\beta$  biomarkers, but time-to-dementia should be predicted by neurodegenerative severity (atrophy on MRI). Our data are consistent with this model in that amyloid positive subjects were far more likely to progress within 2 years than amyloid negative subjects (50 versus 19%). While A $\beta$  deposition initiates the pathological cascade, it is not the direct cause of cognitive impairment. Accordingly, in our data A $\beta$  load severity is decoupled from time-to-progression at high levels. In contrast, hippocampal atrophy indicates how far along the neurodegenerative path one is, and hence how close to progressing to dementia (DeKosky and Scheff, 1990; Terry *et al.*, 1991; Jack *et al.*, 2009; Mormino *et al.*, 2009; Savva *et al.*, 2009), which again matches our data. Moreover, the log relative hazard decreased linearly with increasing hippocampal atrophy, indicating a direct relationship between severity of atrophy and time-to-dementia. Figure 3A is a simplification of the 'biomarker cascade model' from Jack *et al.* (2010) with only the A $\beta$  load and MRI neurodegenerative biomarkers displayed. The 'biomarker cascade model' describes the cognitive and biomarker trajectory of an individual who develops Alzheimer's disease dementia over his/her adult lifetime. In contrast, all subjects in this paper began the study having already demonstrated a cognitive impairment, which indicates that the process of neurodegeneration had already begun. Over the time a subject traverses the horizontal 'clinical distance' (indicated by the horizontal red and blue arrows in Fig. 3A) from the time-of-diagnosis of mild cognitive impairment to a time-of-diagnosis of dementia (indicated by vertical lines in Fig. 3A), the vertical 'distance travelled' along the A $\beta$  load biomarkers curve is small. In contrast, over this same 'clinical distance travelled' on the horizontal axis, the vertical distance travelled along the MRI biomarker is substantial. Thus the position of a subject with mild cognitive impairment along the MRI curve is a strong determinant of his/her time to dementia. One important caveat is that follow-up times in our sample, as in most published studies on mild cognitive impairment, were relatively short, which limits our conclusions as to the

relationships between brain A $\beta$  load, brain atrophy and the risk of progressing over a short interval of time (several years). Conclusions related to longer term risk associated with these two classes of Alzheimer's disease biomarkers will require lengthier follow-up.

## Among all subjects with mild cognitive impairment, combined hippocampal atrophy and brain A $\beta$ load predict progression with comparable power

An interesting feature of our data is that the results differ when subjects with mild cognitive impairment are split into amyloid positive or negative, versus when all subjects with mild cognitive impairment are examined together. When all subjects with mild cognitive impairment are analysed together, hippocampal atrophy and brain A $\beta$  load predict progression to dementia with comparable discriminative power. However, the risk profile associated with increasing severity of these two classes of biomarker differs notably. The log relative hazard is linear throughout the range of hippocampal atrophy values, which means that a unit decrease in hippocampal volume has the same increase in relative hazard (i.e. shortened time-to-progression) across the spectrum of hippocampal values. In contrast, for brain A $\beta$  load, the log relative hazard is approximately linear from 1.0 through  $\sim$ 2.0, at which point it plateaus. This suggests that up to a value of  $\sim$ 2.0 increasing A $\beta$  confers commensurate shortened time-to-progression but values beyond this threshold confer little or no additional apparent relative hazard.

While there are several possible explanations for this finding, one is the possible effect of coexistent pathologies. Mild cognitive impairment is an aetiologically heterogeneous clinical diagnosis. A proportion of subjects in most longitudinal cohorts of subjects with mild cognitive impairment do not progress over long-term clinical follow-up and are felt most likely to have non-progressive conditions (DeCarli, 2003). Among subjects with mild cognitive impairment who progressed from a clinical diagnosis of mild cognitive impairment to dementia and then came to autopsy, Jicha *et al.* (2006) found that 71% had typical Alzheimer's pathology. However, 29% had a final primary neuropathological diagnosis other than Alzheimer's disease, including Lewy body disease, hippocampal sclerosis, non-specific tauopathy, fronto-temporal lobar dementia and cerebrovascular disease. In addition, 82% had two or more pathological processes that were felt to contribute to dementia, including 35% with cerebrovascular disease and 26% with Lewy body disease. Similar autopsy findings in mild cognitive impairment have been reported by others (Bennett *et al.*, 2005a; Markesbery *et al.*, 2006; Schneider *et al.*, 2009). The data in our mild cognitive impairment combined group ( $n=218$ ) can therefore be interpreted from the perspective of certain aetiological heterogeneity. We can safely assume that our total mild cognitive impairment group contained: (i) subjects with Alzheimer's pathology only; (ii) subjects with no Alzheimer's pathology (one of our subjects did progress to semantic dementia and another to multi system atrophy, both diagnosed clinically); (iii) subjects with mild



**Figure 3** Hypothetical effects of pure Alzheimer's pathology versus mixed pathology on time-to-progression from mild cognitive impairment to dementia. Clinical disease stage is indicated on the horizontal axis with vertical lines indicating the time at which diagnoses of mild cognitive impairment and dementia are reached. The severity of A $\beta$  load (red curve) and brain atrophy (blue curve) on the vertical axis range from normal to maximally abnormal. (A) [modified from Jack *et al.* (2010)], illustrates the hypothetical biomarker curves of Subject A, who progresses from normal to mild cognitive impairment to dementia and who has pure Alzheimer's pathology. Over the time a subject traverses the horizontal 'clinical distance' from time-of diagnosis of mild cognitive impairment to time-of diagnosis of dementia (indicated by the horizontal red and blue arrows), the vertical 'distance travelled' along the A $\beta$  load biomarkers curve is small as indicated by the red vertical arrow in (A). In contrast, over this same 'clinical distance travelled' on the horizontal axis, the vertical distance travelled along the MRI biomarker curve is substantial as indicated by the blue vertical arrow in (A). (B) illustrates the hypothetical curve of Subject B, who has mixed pathology. The effect of the coexistent second pathology is to shift the blue atrophy curve, time-of diagnosis of mild cognitive impairment, and time-of diagnosis of dementia closer to the A $\beta$  curve. Consequently Subject B reaches a diagnosis of dementia with a lower level of A $\beta$  *in situ* than Subject A. AD = Alzheimer's disease; MCI = mild cognitive impairment.

cognitive impairment with mixed pathology, particularly Alzheimer's plus cerebrovascular disease or Lewy body disease. While the autopsy numbers at this point are small, to date 10 of our subjects have come to autopsy bearing the clinical diagnosis of either mild cognitive impairment ( $n=2$ ) or Alzheimer's

dementia ( $n=8$ ) at the time of death. Of these, four (40%) had autopsy findings of Alzheimer's pathology without any other significant co-morbidity and six (60%) had autopsy findings of mixed Alzheimer's pathology (Cairns *et al.*, 2010). While subjects with mixed pathology are not addressed by the 'biomarker



cascade model' (Jack *et al.*, 2010), in the next paragraph we expand this model to incorporate the hypothetical effect of coexistent pathologies.

To interpret our results more easily in light of the 'biomarker cascade model' (Jack *et al.*, 2010), we divided our amyloid positive group ( $n=165$ ) into 113 subjects with high A $\beta$  load ( $>2.0$ ) and 52 subjects with intermediate A $\beta$  load (1.5–2.0). Subjects with mild cognitive impairment and intermediate A $\beta$  load levels do not completely fit the 'biomarker cascade model', because the model predicts that by the time a subject is sufficiently impaired to reach a clinical diagnosis of mild cognitive impairment, he/she already has accumulated a substantial A $\beta$  load. There are several possible explanations for this: some subjects may reach a plateau at lower absolute levels of A $\beta$  load, some subjects may be more biologically susceptible to A $\beta$  or levels of soluble A $\beta$  oligomers may be more critical than fibrillar A $\beta$  levels. One interesting possibility, however, is that subjects with mild cognitive impairment and intermediate levels of A $\beta$  probably represent individuals with mixed Alzheimer's-other pathology. The proportion of APOE  $\epsilon 4$  carriers among the subjects with high A $\beta$  load and mild cognitive impairment was 84%, but only 29% among the subjects with intermediate A $\beta$  load and mild cognitive impairment. This supports the idea that those with intermediate A $\beta$  load values in our sample were more likely to have co-morbid non-Alzheimer's pathologies contributing to cognitive impairment than the subjects with high A $\beta$  load. Cerebrovascular disease is the second most common pathological substrate associated with dementia in most autopsy series (Schneider *et al.*, 2004, 2007a; Sonnen *et al.*, 2007). Recent work suggests that micro infarction in particular is a key cerebrovascular disease pathology that contributes to cognitive impairment (Petrovitch *et al.*, 2005; White *et al.*, 2005; Sonnen *et al.*, 2007; Schneider *et al.*, 2009). The implication is that cerebrovascular disease is a separate pathway to degenerative brain atrophy (Jagust *et al.*, 2008) that has an additive effect on cognition along with the neurodegeneration initiated by amyloidosis (Schneider *et al.*, 2007b, 2009; Jagust *et al.*, 2008). With both amyloidogenic and cerebrovascular disease pathologies in operation, an individual can receive a diagnosis of mild cognitive impairment with a lower level of amyloid pathology than an individual with only amyloid-initiated degenerative pathology. We illustrate this effect hypothetically in Fig. 3. Whereas Fig. 3A shows a subject with pure Alzheimer's pathology, Fig. 3B shows a hypothetical subject with mixed Alzheimer's-cerebrovascular disease pathology. Hypothetical curves for A $\beta$  load and hippocampal atrophy are indicated as non-linear functions of time and are indexed to the points at which diagnoses of mild cognitive impairment and dementia are made (indicated by vertical lines). We model the effect of combined pathology in Fig. 3B as a decrease in the temporal separation between the A $\beta$  amyloid curve and the complex comprised of the brain atrophy curve and the vertical markers indicating time-of-mild cognitive impairment and dementia diagnoses. The effect is that Subject B becomes demented with lower levels of A $\beta$  load than Subject A, however, the linear relationship between brain atrophy and time to dementia is the same for both subjects. Note that in Fig. 3B we do not intend to imply that subjects with mixed pathology necessarily

reach a diagnosis of dementia at a younger age, but rather do so with a lower A $\beta$  level than subjects with pure Alzheimer's pathology. We acknowledge that this explanation is speculative because we cannot know the status of co-existent pathologies in individual subjects *ante mortem*. However, there is no reason to suspect that the subjects with mild cognitive impairment enrolled in the Alzheimer's Disease Neuroimaging Initiative are notably different from subjects with mild cognitive impairment in earlier autopsy studies where the prevalence of coexistent pathologies has been consistently well documented.

## 'Brain A $\beta$ load' measurement: converting cerebrospinal fluid A $\beta$ 42 measures to Pittsburgh compound B positron emission tomography units and combining measures across subjects

A $\beta$  load was ascertained by CSF in 165 subjects and by PIB PET imaging in 53 subjects. We recognize that using only global A $\beta$  load values ignores potentially useful regional information in PIB PET images. We also emphasize the limitations inherent in pooling CSF and PIB PET data. However, pooling subjects with either CSF or PIB PET increased our sample size and improved statistical power by increasing the number of progression events. In our case, the increased sample size was necessary to adequately power the evaluation of non-linearity in the MRI and A $\beta$  load risk profiles. Justification for pooling CSF and imaging-based measurements is the consistently observed tight inverse correlation between PIB PET and CSF A $\beta$ 42 measures in every study where the two measures have been compared (Fagan *et al.*, 2006; Forsberg *et al.*, 2008; Grimmer *et al.*, 2009; Jagust *et al.*, 2009; Degerman Gunnarsson *et al.*, 2010). We used the multiple imputation measurement error (Cole *et al.*, 2006) method of transforming CSF A $\beta$ 42 into PIB PET units to produce statistically equivalent measures of A $\beta$  load from either biomarker source (Weigand *et al.*, 2010). The multiple imputation measurement error method assumes that two measures of the same biological phenomenon exist. One is the 'gold standard'; the second is a surrogate measurement with associated measurement error. We considered PIB PET imaging to be the more direct measure of A $\beta$  load in the brain and CSF-based A $\beta$ 42 to be a surrogate measured with additional error. Our approach can therefore be thought of as calibrating or adjusting the findings from CSF to the more accurate imaging based measure while taking into account uncertainty in the calibration process.

One concern we had was that the non-linear log relative hazard of A $\beta$  load illustrated in Fig. 2 could be an artefact of the model transforming CSF A $\beta$ 42 into PIB PET units. To address this, we also performed the time-to-event analysis on the rank-transformed A $\beta$  load. Our findings of non-linearity of A $\beta$  load and linearity of hippocampal volume did not change using this rank-based analysis. Because the rank-transformed values would be the same under any reasonable progression model, we concluded that the non-linearity in the time-to-event data reflects a real biological property of A $\beta$  load and not a feature of the relationship

between PIB PET and CSF A $\beta$ 42. In addition, Fig. 2 illustrates that *APOE*  $\epsilon$ 4 carriers are more likely to have higher A $\beta$  load than non-carriers. That is, the established property of *APOE*  $\epsilon$ 4 as primarily a risk factor for brain A $\beta$  deposition (Schmechel *et al.*, 1993; Morris *et al.*, 2010; Vemuri *et al.*, 2010) is maintained when the A $\beta$  load is measured from pooling CSF and PIB PET data.

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